

Advanced Bayesian Methods: Theory and Applications in R

Big Data and Variable Selection

Nikolaus Umlauf https://nikum.org/abm.html

Estimation

The main building block of regression model algorithms is the probability density function $d_y(\mathbf{y}|\theta_1, \ldots, \theta_K)$.

Estimation typically requires to evaluate the log-likelihood

$$\ell(\boldsymbol{\beta}; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \log d_{y}(y_{i}; \theta_{1}(\mathbf{x}_{i}; \boldsymbol{\beta}_{1}), \ldots, \theta_{K}(\mathbf{x}_{i}; \boldsymbol{\beta}_{K})),$$

with $\mathbf{X} = (\mathbf{X}_1, \dots, \mathbf{X}_K)$.

The log-posterior (frequentist penalized log-likelihood)

$$\log \pi(eta, au; \mathbf{y}, \mathbf{X}, lpha) \propto \ell(eta; \mathbf{y}, \mathbf{X}) + \sum_{k=1}^{K} \sum_{j=1}^{J_k} \left[\log p_{jk}(eta_{jk}; au_{jk}, lpha_{jk})
ight],$$

where $p_{jk}(\cdot)$ are priors, τ_{jk} (smoothing) variances and α_{jk} fixed hyper parameters.

Priors
$$p_{jk}(\cdot)$$

For simple linear effects $\mathbf{X}_{jk}\beta_{jk}$: $p_{jk}(\beta_{jk}) \propto const$.

For the smooth terms:

$$p_{jk}(eta_{jk}; au_{jk},lpha_{jk}) \propto d_{eta_{jk}}(eta_{jk}|\, au_{jk}; lpha_{eta_{jk}}) \cdot d_{ au_{jk}}(au_{jk}|\, lpha_{ au_{jk}}).$$

Using a basis function approach a common choice is

$$d_{eta_{jk}}(eta_{jk}|\,m{ au}_{jk},lpha_{eta_{jk}}) \propto |m{ extbf{P}}_{jk}(m{ au}_{jk})|^rac{1}{2} \exp\left(-rac{1}{2}m{eta}_{jk}^ opm{ extbf{P}}_{jk}(m{ au}_{jk})m{eta}_{jk}
ight).$$

Precision matrix $\mathbf{P}_{jk}(\tau_{jk})$ derived from prespecified penalty matrices $\alpha_{\beta_{jk}} = {\mathbf{K}_{1jk}, \dots, \mathbf{K}_{Ljk}}.$

The variances parameters τ_{jk} are equivalent to the inverse smoothing parameters in a frequentist approach.

Estimation

Bayesian point estimates of parameters are obtained by:

- Maximization of the log-posterior for posterior mode estimation.
- Solving high dimensional integrals, e.g., for posterior mean or median estimation.

Problems 1 and 2 are commonly solved by computer intensive iterative algorithms of the following type:



Typically the number of different observations $x_{(1)} < x_{(2)} < \cdots < x_{(m)}$ in a design matrix **X** is much smaller than the total number *n* of observations, i.e., $m \ll n$. For **sorted** observations x_i :

- Index vector ind with $ind[i] \in \{1, \ldots, m\}$, i.e., if $x_i = x_{(s)}$ then ind[i] = s.
- Decompose the design matrix in $\textbf{X} = \textbf{DP}\tilde{\textbf{X}}$ where
- X is the m × L reduced design matrix for the different and sorted observations x₍₁₎,..., x_(m), i.e., X [s, l] = X_l(x_s), s = 1,...,m, l = 1,...,L,
- **P** is a $n \times L$ permutation matrix, which reverts the sorting, i.e., **P**[*i*, *s*] = *l*(**ind**[*i*] = *s*).
- **D** is a diagonal matrix, e.g., for varying coefficient models or **D** = **I** for simple additive terms.
- For the vector of function evaluations we obtain $\mathbf{f} = \mathbf{X}\boldsymbol{\beta} = \mathbf{D}\mathbf{P}\tilde{\mathbf{X}}\boldsymbol{\beta}$.

Using the permutation, we get

$$\mathbf{X}_{jk}^{\top}\mathbf{W}_{kk}\mathbf{X}_{jk} = \mathbf{\tilde{X}}_{jk}^{\top}\mathbf{P}_{jk}^{\top}\mathbf{D}_{jk}^{\top}\mathbf{W}_{kk}\mathbf{D}_{jk}\mathbf{P}_{jk}\mathbf{\tilde{X}}_{jk} = \mathbf{\tilde{X}}_{jk}^{\top}\mathbf{\tilde{W}}\mathbf{\tilde{X}}_{jk},$$

where

$$\tilde{\mathbf{W}} = \mathbf{P}_{jk}^{\top} \mathbf{D}_{jk}^{\top} \mathbf{W}_{kk} \mathbf{D}_{jk} \mathbf{P}_{jk} = diag(\tilde{w}_1, \dots, \tilde{w}_{m_{jk}})$$

and the "reduced" weights \tilde{w}_s , are given by

$$\tilde{w}_s = \sum_{i: ind[i]=s} z_i^2 \mathbf{W}_{kk}[i,i].$$

The weights \tilde{w}_s can be computed by first initializing $\tilde{w}_s = 0$ followed by a simple loop:

For
$$i = 1, ..., n$$
 add $z_i^2 \mathbf{W}_{kk}[i, i]$ to $\tilde{w}_{ind[i]}$.

For $\mathbf{X}_{jk}^{ op} \mathbf{W}_{kk}(\mathbf{z}_k - m{\eta}_{k,-j}^{(t)})$ we obtain

$$\mathbf{X}_{jk}^{\top}\mathbf{W}_{kk}\mathbf{r} = \tilde{\mathbf{X}}_{jk}^{\top}\mathbf{P}_{jk}^{\top}\mathbf{D}_{jk}^{\top}\mathbf{W}_{kk}\mathbf{r} = \tilde{\mathbf{X}}_{jk}^{\top}\tilde{\mathbf{r}},$$

with partial residuals $\mathbf{r} = \mathbf{z}_k - \eta_{k,-j}^{(t)}$.

The "reduced" partial residuals yield a $m_{jk} imes 1$ vector $\tilde{\mathbf{r}} = (\tilde{r}_1, \dots, \tilde{r}_{m_{jk}})^\top$ given by

$$\tilde{r}_s = \sum_{i: \operatorname{ind}[i]=s} z_i \mathbf{W}_{kk}[i,i] r_i.$$

The \tilde{r}_s are computed by first initializing $\tilde{r}_s = 0$ followed by the loop: For i = 1, ..., n add $z_i \mathbf{W}_{kk}[i, i] r_i$ to \tilde{r}_{indial} .

Example using the IndiaNutrition data set.

```
R> dim(IndiaNutrition)
[1] 25134 13
R> X <- smoothCon(s(mage, bs = "ps", k = 22),
+ IndiaNutrition. NULL)[[1]]$X</pre>
```

```
+ IndiaNutrition, NULL)[[1]]$X
R> dim(X)
[1] 25134 22
R> i <- match.index(X)
R> tX <- X[i$nodups, ]
R> dim(tX)
[1] 26 22
R> print(object.size(X), units = "Mb")
4.2 Mb
```

```
R> print(object.size(tX), units = "Kb")
```

4.7 Kb

Sparsity

B-spline design matrix:

$\mathbf{X}_{jk} =$	(0.496	0.504	0	0	0	0
	0.153	0.747	0.1	0	0	0
	0.006	0.597	0.397	0	0	0
	0	0.221	0.723	0.056	0	0
	0	0.025	0.673	0.303	0	0
	0	0	0.303	0.673	0.025	0
	0	0	0.056	0.723	0.221	0
	0	0	0	0.397	0.597	0.006
	0	0	0	0.1	0.747	0.153
	0 /	0	0	0	0.504	0.496

Sparsity

B-spline penalty matrix:

$$\mathbf{K}_{jk} = \begin{pmatrix} 0.25 & -0.25 & 0 & 0 & 0 & 0 \\ -0.25 & 0.5 & -0.25 & 0 & 0 & 0 \\ 0 & -0.25 & 0.5 & -0.25 & 0 & 0 \\ 0 & 0 & -0.25 & 0.5 & -0.25 & 0 \\ 0 & 0 & 0 & -0.25 & 0.5 & -0.25 \\ 0 & 0 & 0 & 0 & -0.25 & 0.25 \end{pmatrix}$$



Markov random fields (MRF) design matrix:

Products $\tilde{\mathbf{X}}_{jk}^{\top} \tilde{\mathbf{W}} \tilde{\mathbf{X}}_{jk}$ and $\tilde{\mathbf{X}}_{jk}^{\top} \tilde{\mathbf{r}}$ are stored in sparse matrix format. Nonzero entries are stored in a vector **C** ($n_x \times 1$). E.g., the *t*-th entry **C**[*t*] corresponds to

$$\mathbf{C}[t] = \sum_{s=1}^{m_{jk}} \tilde{w}_s \tilde{\mathbf{X}}_{jk}[s, r] \tilde{\mathbf{X}}_{jk}[s, l],$$

hence, most products are zero. Store the nonzero products in \mathbf{h}_1 , the nonzero index *s* in \mathbf{h}_2 and the position of the first element in \mathbf{h}_3 . Computation only requires

$$\mathbf{C}[t] = \sum_{s=\mathbf{h}_3[t]}^{\mathbf{h}_3[t+1]-1} \tilde{w}_{\mathbf{h}_2[s]} \mathbf{h}_1[s].$$

Similarly for $\tilde{\mathbf{X}}_{ik}^{\top}\tilde{\mathbf{r}}$, etc.

Example using the IndiaNutrition data set.
R> H <- sparse.matrix.index(tX)
R> print(head(H))

```
[,1] [,2] [,3] [,4]
[1.] 4 5 6
[2,] 3 4 5 6
[3,] 2 3 4 5
[4.] 9 10 11 12
[5,] 4 5 6 7
[6.]
          6 7
      5
                  8
R> print(nrow(X) * ncol(X))
[1] 552948
R> print(nrow(tX) * ncol(tX))
[1] 572
R> print(nrow(H) * ncol(H))
[1] 104
R> print(object.size(H), units = "Kb")
0.6 Kb
```

Consider the following updating scheme

$$eta_k^{[t+1]} = \mathtt{U}_k(eta_k^{[t]}; \cdot) = eta_k^{[t]} - \mathtt{H}_{kk}\left(eta_k^{[t]}
ight)^{-1} \mathtt{s}\left(eta_k^{[t]}
ight).$$

Assuming model terms that can be written as a matrix product of a design matrix and coefficients we obtain an iteratively weighted least squares scheme given by

$$\beta_{jk}^{[t+1]} = \mathtt{U}_{jk}(\beta_{jk}^{[t]}; \cdot) = (\mathbf{X}_{jk}^{\top} \mathbf{W}_{kk} \mathbf{X}_{jk} + \mathbf{G}_{jk}(\boldsymbol{\tau}_{jk}))^{-1} \mathbf{X}_{jk}^{\top} \mathbf{W}_{kk}(\mathbf{z}_k - \boldsymbol{\eta}_{k,-j}^{[t+1]}),$$

with working observations $\mathbf{z}_k = \eta_k^{[t]} + \mathbf{W}_{kk}^{-1[t]} \mathbf{u}_k^{[t]}$, working weights $\mathbf{W}_{kk}^{-1[t]}$ and score vector $\mathbf{u}_k^{[t]}$.

Instead of using all observations of the data, we only use a randomly chosen **subset** denoted by the subindex [**s**] in one updating step

$$\begin{split} \boldsymbol{\beta}_{jk}^{[t+1]} &= \nu \cdot (\mathbf{X}_{[\mathbf{s}],jk}^{\top} \mathbf{W}_{[\mathbf{s}],kk} \mathbf{X}_{[\mathbf{s}],jk} + \mathbf{G}_{jk}(\boldsymbol{\tau}_{jk}))^{-1} \mathbf{X}_{[\mathbf{s}],jk}^{\top} \mathbf{W}_{[\mathbf{s}],kk} (\mathbf{z}_{[\mathbf{s}],k} - \eta_{[\mathbf{s}],k,-j}^{[t+1]}) + \\ & (1-\nu) \cdot \boldsymbol{\beta}_{jk}^{[t]}, \end{split}$$

where ν is a weight parameter which specifies how much the parameters at iteration t + 1 are influenced by parameters of the previous iteration t.

Use **flat file** format for each X_{jk} , i.e., only batch [s] is in memory. This way, we can estimate models with **really** large datasets!

Mimics a second order stochastic gradient descent (SGD) algorithm

$$eta_{jk}^{[t+1]} = eta_{jk}^{[t]} +
u \cdot (eta_{jk,[\mathbf{s}]} - eta_{jk}^{[t]}) = eta_{jk}^{[t]} +
u \cdot eta_{jk}^{[t]},$$

and $\delta_{ik}^{[t]}$ is composed from first and second order derivative information with

$$\begin{split} \boldsymbol{\delta}_{jk}^{[t]} &= \boldsymbol{\beta}_{jk,[\mathbf{s}]} - \boldsymbol{\beta}_{jk}^{[t]} \\ &= \left[\boldsymbol{\beta}_{jk}^{[t]} - \mathbf{H}_{[\mathbf{s}],kk} \left(\boldsymbol{\beta}_{jk}^{[t]} \right)^{-1} \mathbf{s}_{[\mathbf{s}]} \left(\boldsymbol{\beta}_{jk}^{[t]} \right) \right] - \boldsymbol{\beta}_{jk}^{[t]} \\ &= -\mathbf{H}_{[\mathbf{s}],kk} \left(\boldsymbol{\beta}_{jk}^{[t]} \right)^{-1} \mathbf{s}_{[\mathbf{s}]} \left(\boldsymbol{\beta}_{jk}^{[t]} \right) \end{split}$$

Hence, the updating step length is adaptive.

The idea is to select τ_{jk} using a stepwise algorithm which is based on an **"out-of-sample" criterion**, i.e., the criterion $C(\cdot)$ is evaluated on another batch denoted by $[\tilde{\mathbf{s}}]$, $C_{[\tilde{\mathbf{s}}]}(\cdot)$ respectively, i.e.

$$\tau_{ljk}^{[t+1]} \leftarrow \operatorname*{arg\,min}_{\tau_{ljk}^{\star} \in \mathcal{I}_{ljk}} C_{[\mathbf{\tilde{s}}]}(U_{jk}(\beta_{jk}^{[t]}, \tau_{ljk}^{\star}; \cdot)),$$

where \mathcal{I}_{ljk} is a search interval for $\tau_{ljk}^{[t+1]}$, e.g.,

$$\mathcal{I}_{ljk} = [\tau_{ljk}^{[t]} \cdot \mathbf{10^{-1}}, \tau_{ljk}^{[t]} \cdot \mathbf{10}].$$

Some interesting features:

- **1** Set, e.g., $\nu = 0.1$, convergence after visiting *m* batches **[s**].
- Only update if "out-of-sample" log-likelihood is increased.
- **Boosting** for **variable selection**: Update only $f_{jk}(\cdot)$ with greatest contribution in "out-of-sample" log-likelihood.
- **3 Bagging**: If $\nu = 1$, each update is so to say a "**sample**". Convergence similar to MCMC algorithms, i.e., if $\beta_{jk}^{[t+1]}$ start fluctuating around a certain level.
- **5** Slice sample τ_{ljk} under $C_{[\tilde{s}]}(\cdot)$, much faster!

- Project aiming to better explain the problems of childhood malnutrition in low- and middle-income countries.
- Contribute to monitoring of the Sustainable Development Goals (SGD).
- We compiled a brand new data set using DHS data.
- Data on global conflicts, topography and environmental data from satellite observations (NDVI), temperature and precipitation data from ERA5 (ECMWF).
- Data from 1990–2019 with n > 3M observations.

Application



Application





Example: Search distribution.

Define the batchsize.

R> bs <- 2000

Generate batches.

```
R> batch_ids <- lapply(1:200, function(...) {
+ sample(1:nrow(d), size = bs, replace = FALSE)
+ })</pre>
```

Estimate model.

```
R> b <- bamlss(y ~ 1, data = d, family = JSU,
+ sampler = FALSE, optimizer = opt_bbfitp, slice = 10, aic = TRUE, K = 2,
+ batch_ids = batch_ids)
```

Compute log-likelihood.

R> logLik(b, newdata = nd)



Example: Boosting flavour with *ff* data frame.

Set up a model formula.

```
R> f <- list(
+ stunting ~ s(cage) + s(bord) + s(hhs) + s(x, y) + ...,
+ ~ ~ s(cage) + s(bord) + s(hhs) + s(x, y) + ...
+ )</pre>
```

Estimate model.

```
R> b <- bamlss(f, data = dff, family = JSU,
+ sampler = FALSE, optimizer = opt_bbfit,
+ batch_ids = batch_ids, select = TRUE, aic = TRUE, always = FALSE,
+ eps_loglik = 0.001, K = 2, overwrite = TRUE, delete = FALSE,
+ ff_name = ff_name)
```

Plot results.

R> contribplot(b)



Example: Bagging type flavour with slice sampling.

Extract formula.

```
R> nf <- new_formula(b)
```

Estimate model using ff processed data.

```
R> m <- bamlss(nf, data = dff, family = JSU,
+ sampler = FALSE, optimizer = opt_bbfitp,
+ batch_ids = batch_ids, aic = TRUE, slice = TRUE,
+ ff_name = ff_name)
```

Afterwards, all extractor functions provided by *bamlss* can be used, e.g., summary(), predict(), plot(), etc.

R> plot(m, which = "samples")

